

Bayesian Linear Regression

Guest lecture at KTH 2020

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Lecture overview

- Bayesian inference
- The normal model with known variance
- Linear regression
- Regularization priors

Slides at: <https://mattiasvillani.com/news>

Likelihood function - normal data regression

- Normal data with known variance:

$$X_1, \dots, X_n | \theta \stackrel{iid}{\sim} \mathcal{N}(\theta, \sigma^2).$$

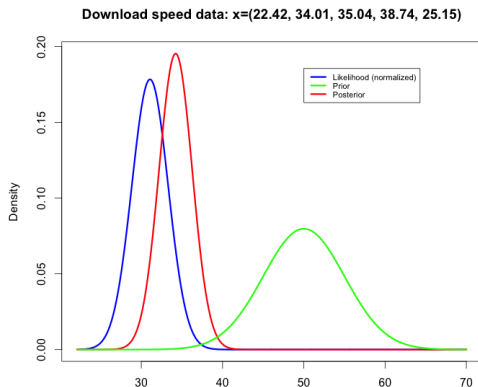
- Likelihood from independent observations: x_1, \dots, x_n

$$\begin{aligned} p(x_1, \dots, x_n | \theta) &= \prod_{i=1}^n p(x_i | \theta) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2\right) \\ &\propto \exp\left(-\frac{1}{2(\sigma^2/n)} (\theta - \bar{x})^2\right) \end{aligned}$$

- Maximum likelihood: $\hat{\theta} = \bar{x}$ maximizes $p(x_1, \dots, x_n | \theta)$.
- Given the data x_1, \dots, x_n , plot $p(x_1, \dots, x_n | \theta)$ as a function of θ .

Am I really getting my 50Mbit/sec?

- My broadband provider promises me at least 50Mbit/sec.
- **Data:** $x = (22.42, 34.01, 35.04, 38.74, 25.15)$ Mbit/sec.
- **Measurement errors:** $\sigma = 5$ (± 10 Mbit with 95% probability)
- The likelihood function is proportional to $N(\bar{x}, \sigma^2/n)$ density.



The likelihood function

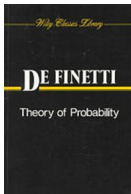
- The mantra:

*The likelihood function is
the probability of the observed data
considered as a function of the parameter.*

- Likelihood function is **NOT** a probability distribution for θ .
- Statements like $\Pr(\theta \geq 50 | \text{data})$ makes no sense.
- Unless ...

Uncertainty and subjective probability

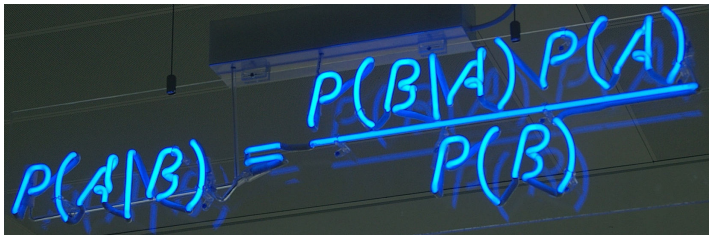
- $\Pr(\theta \geq 50|\text{data})$ only makes sense if θ is random.
- But θ may be a fixed natural constant?
- **Bayesian: doesn't matter if θ is fixed or random.**
- Do **You** know the value of θ or not?
- $p(\theta)$ reflects Your knowledge/**uncertainty** about θ .
- **Subjective probability.**
- The statement $\Pr(\text{10th decimal of } \pi = 9) = 0.1$ makes sense.



Bayesian learning

- **Bayesian learning** about a model parameter θ :
 - ▶ state your **prior** knowledge as a probability distribution $p(\theta)$.
 - ▶ collect **data** \mathbf{x} and form the **likelihood** function $p(\mathbf{x}|\theta)$.
 - ▶ **combine** prior knowledge $p(\theta)$ with data information $p(\mathbf{x}|\theta)$.
- **How to combine** the two sources of information?

Bayes' theorem



A photograph of a whiteboard with the equation for Bayes' theorem written in blue marker. The equation is
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

Learning from data - Bayes' theorem

- How to **update** from **prior** $p(\theta)$ to **posterior** $p(\theta|Data)$?
- **Bayes' theorem** for events A and B

$$p(A|B) = \frac{p(B|A)p(A)}{p(B)}.$$

- Bayes' Theorem for a model parameter θ

$$p(\theta|Data) = \frac{p(Data|\theta)p(\theta)}{p(Data)}.$$

- It is the prior $p(\theta)$ that takes us from $p(Data|\theta)$ to $p(\theta|Data)$.
- A probability distribution for θ is extremely useful:
 - ▶ **Predictions**
 - ▶ **Decision making**
 - ▶ **Regularization**

Great theorems make great tattoos

- Bayes theorem

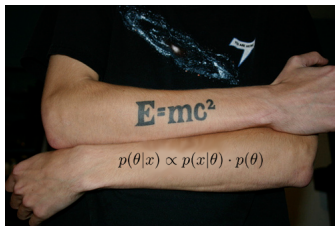
$$p(\theta|Data) = \frac{p(Data|\theta)p(\theta)}{p(Data)}$$

- All you need to know:

$$p(\theta|Data) \propto p(Data|\theta)p(\theta)$$

or

$$\text{Posterior} \propto \text{Likelihood} \cdot \text{Prior}$$



Normal data, known variance - uniform prior

■ Model

$$x_1, \dots, x_n | \theta, \sigma^2 \stackrel{iid}{\sim} N(\theta, \sigma^2).$$

■ Prior

$$p(\theta) \propto c \text{ (a constant)}$$

■ Likelihood

$$p(x_1, \dots, x_n | \theta, \sigma^2) = \exp \left[-\frac{1}{2(\sigma^2/n)} (\theta - \bar{x})^2 \right]$$

■ Posterior

$$\theta | x_1, \dots, x_n \sim N(\bar{x}, \sigma^2/n)$$

Normal data, known variance - normal prior

■ Prior

$$\theta \sim N(\mu_0, \tau_0^2)$$

■ Posterior

$$\begin{aligned} p(\theta | x_1, \dots, x_n) &\propto p(x_1, \dots, x_n | \theta, \sigma^2) p(\theta) \\ &\propto N(\theta | \mu_n, \tau_n^2), \end{aligned}$$

where

$$\frac{1}{\tau_n^2} = \frac{n}{\sigma^2} + \frac{1}{\tau_0^2},$$

$$\mu_n = w\bar{x} + (1 - w)\mu_0,$$

and

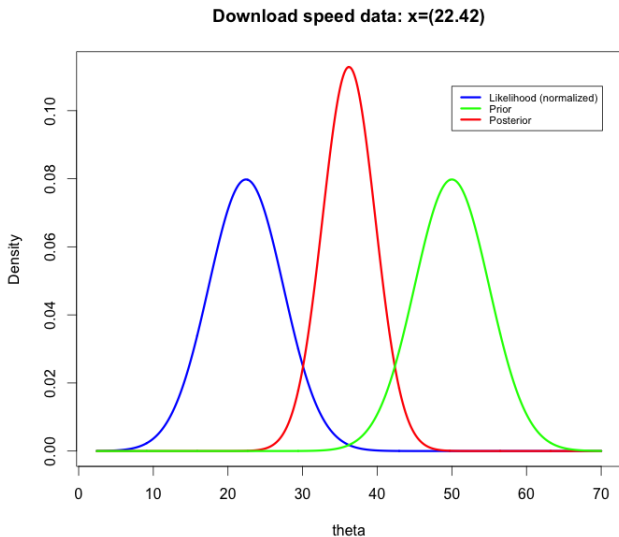
$$w = \frac{\frac{n}{\sigma^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}}.$$

■ Proof: complete the squares in the exponential.

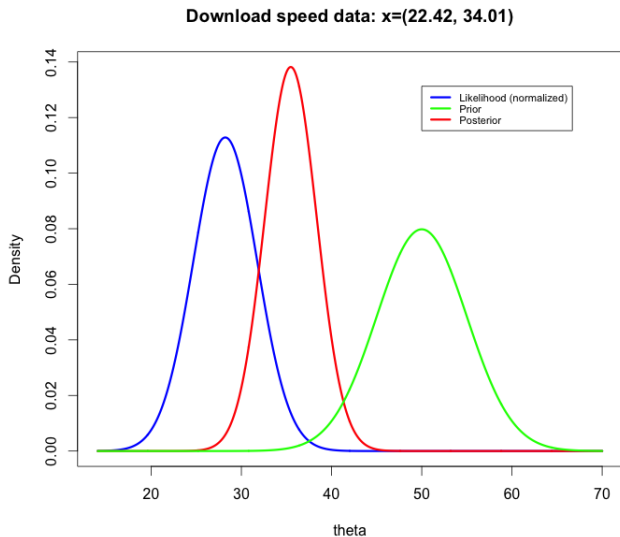
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- **Data:** $x = (22.42, 34.01, 35.04, 38.74, 25.15)$ Mbit/sec.
- **Model:** $X_1, \dots, X_5 \sim N(\theta, \sigma^2)$.
- Assume $\sigma = 5$ (measurements can vary ± 10 MBit with 95% probability)
- My **prior:** $\theta \sim N(50, 5^2)$.

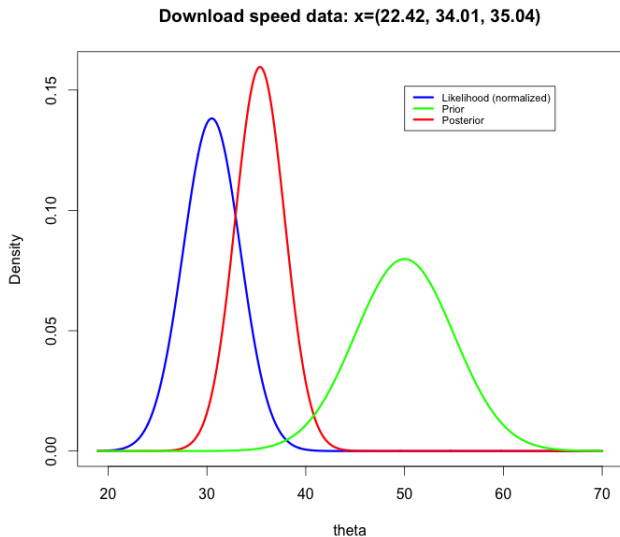
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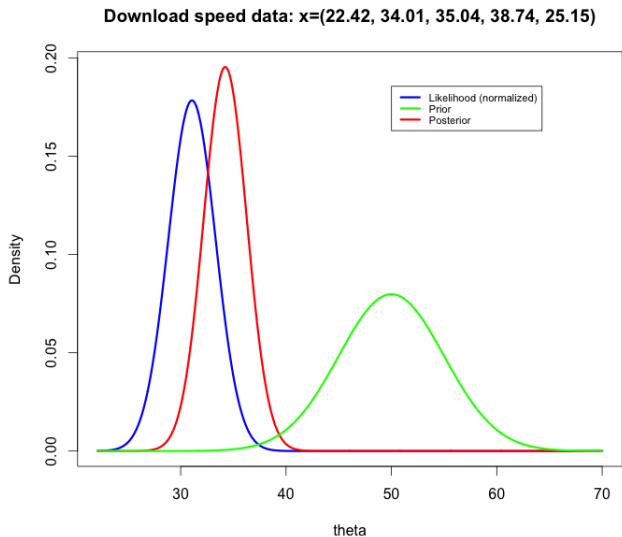
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Download speed $n=3$



Download speed $n=5$



Linear regression

- The linear regression model in **matrix form**

$$\underset{(n \times 1)}{\mathbf{y}} = \underset{(n \times k)(k \times 1)}{\mathbf{X}\boldsymbol{\beta}} + \underset{(n \times 1)}{\boldsymbol{\varepsilon}}$$

- Usually first column of \mathbf{X} is the unit vector and β_1 is the intercept.
- Normal errors: $\varepsilon_j \stackrel{iid}{\sim} N(0, \sigma^2)$, so $\boldsymbol{\varepsilon} \sim N(0, \sigma^2 \mathbf{I}_n)$.

- **Likelihood**

$$\mathbf{y} | \boldsymbol{\beta}, \sigma^2, \mathbf{X} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I}_n)$$

Linear regression - uniform prior

- Standard **non-informative prior**: uniform on $(\beta, \log \sigma^2)$

$$p(\beta, \sigma^2) \propto \sigma^{-2}$$

- **Joint posterior** of β and σ^2 :

$$\begin{aligned}\beta | \sigma^2, \mathbf{y} &\sim N[\hat{\beta}, \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}] \\ \sigma^2 | \mathbf{y} &\sim \text{Inv-}\chi^2(n - k, s^2)\end{aligned}$$

where $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$ and $s^2 = \frac{1}{n-k}(\mathbf{y} - \mathbf{X}\hat{\beta})'(\mathbf{y} - \mathbf{X}\hat{\beta})$.

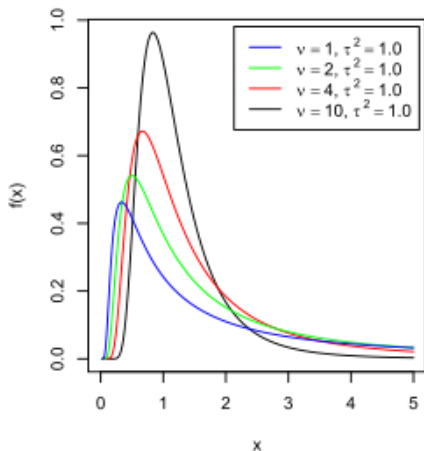
- **Simulate** from the joint posterior by simulating from

- ▶ $p(\sigma^2 | \mathbf{y})$
- ▶ $p(\beta | \sigma^2, \mathbf{y})$

- **Marginal posterior** of β :

$$\beta | \mathbf{y} \sim t_{n-k}[\hat{\beta}, s^2 (\mathbf{X}'\mathbf{X})^{-1}]$$

Scaled inverse χ^2 distribution



- Inverse gamma distribution.

Linear regression - conjugate prior

■ Joint prior for β and σ^2

$$\begin{aligned}\beta|\sigma^2 &\sim N(\mu_0, \sigma^2\Omega_0^{-1}) \\ \sigma^2 &\sim \text{Inv} - \chi^2(v_0, \sigma_0^2)\end{aligned}$$

■ Posterior

$$\begin{aligned}\beta|\sigma^2, \mathbf{y} &\sim N[\mu_n, \sigma^2\Omega_n^{-1}] \\ \sigma^2|\mathbf{y} &\sim \text{Inv} - \chi^2(v_n, \sigma_n^2)\end{aligned}$$

$$\mu_n = (\mathbf{X}'\mathbf{X} + \Omega_0)^{-1} (\mathbf{X}'\mathbf{X}\hat{\beta} + \Omega_0\mu_0)$$

$$\Omega_n = \mathbf{X}'\mathbf{X} + \Omega_0$$

$$v_n = v_0 + n$$

$$v_n\sigma_n^2 = v_0\sigma_0^2 + (\mathbf{y}'\mathbf{y} + \mu_0'\Omega_0\mu_0 - \mu_n'\Omega_n\mu_n)$$

Ridge regression = normal prior

- Problem: too many covariates leads to **over-fitting**.
- **Smoothness/shrinkage/regularization prior**

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} N\left(0, \frac{\sigma^2}{\lambda}\right)$$

- Larger λ gives smoother fit. Note: $\Omega_0 = \lambda I$.
- Equivalent to **penalized likelihood**:

$$-2 \cdot \log p(\beta | \sigma^2, \mathbf{y}, \mathbf{X}) \propto (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta) + \lambda \beta' \beta$$

- Posterior mean gives **ridge regression** estimator

$$\tilde{\beta} = (\mathbf{X}'\mathbf{X} + \lambda I)^{-1} \mathbf{X}'\mathbf{y}$$

- **Shrinkage** toward zero

$$\text{As } \lambda \rightarrow \infty, \tilde{\beta} \rightarrow 0$$

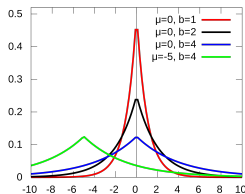
- When $\mathbf{X}'\mathbf{X} = I$

$$\tilde{\beta} = \frac{1}{1 + \lambda} \hat{\beta}$$

Lasso regression = Laplace prior

- Lasso is equivalent to posterior mode under Laplace prior

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} \text{Laplace} \left(0, \frac{\sigma^2}{\lambda} \right)$$



- Laplace prior:
 - ▶ heavy tails
 - ▶ many β_i close to zero, but some β_i can be very large.
- Normal prior
 - ▶ light tails
 - ▶ all β_i 's are similar in magnitude and no β_i very large.

Estimating the shrinkage

- Cross-validation is often used to determine the degree of smoothness, λ .
- Bayesian: λ is **unknown** \Rightarrow **use a prior** for λ .
- $\lambda \sim \text{Inv-}\chi^2(\eta_0, \lambda_0)$. The user specifies η_0 and λ_0 .
- Hierarchical setup:

$$\begin{aligned}\mathbf{y}|\beta, \mathbf{X} &\sim N(\mathbf{X}\beta, \sigma^2 I_n) \\ \beta|\sigma^2, \lambda &\sim N(0, \sigma^2 \lambda^{-1} I_m) \\ \sigma^2 &\sim \text{Inv-}\chi^2(\nu_0, \sigma_0^2) \\ \lambda &\sim \text{Inv-}\chi^2(\eta_0, \lambda_0)\end{aligned}$$

Regression with estimated shrinkage

- The **joint posterior** of β , σ^2 and λ is

$$\beta | \sigma^2, \lambda, \mathbf{y} \sim N(\mu_n, \Omega_n^{-1})$$

$$\sigma^2 | \lambda, \mathbf{y} \sim \text{Inv} - \chi^2(v_n, \sigma_n^2)$$

$$p(\lambda | \mathbf{y}) \propto \sqrt{\frac{|\Omega_0(\lambda)|}{|\mathbf{X}^T \mathbf{X} + \Omega_0(\lambda)|}} \left(\frac{v_n \sigma_n^2(\lambda)}{2} \right)^{-v_n/2} \cdot p(\lambda)$$

- $\Omega_0(\lambda) = \lambda I_m$, and $p(\lambda)$ is the prior for λ .

Polynomial regression

Polynomial regression

$$f(x_i) = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_k x_i^k.$$

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon},$$

where

$$\mathbf{X} = (1, x, x^2, \dots, x^k).$$

- Problem: higher order polynomials can overfit the data.
- Solution: shrink higher order coefficients harder:

$$\boldsymbol{\beta} | \sigma^2 \sim N \left[0, \begin{pmatrix} 100 & 0 & 0 & \dots & 0 \\ 0 & \frac{1}{\lambda} & 0 & \dots & 0 \\ 0 & 0 & \frac{1}{2\lambda} & & \\ \vdots & \vdots & & \ddots & \\ 0 & 0 & 0 & \dots & \frac{1}{k\lambda} \end{pmatrix} \right]$$

Finding the time for maximum

- Quadratic relationship between pain relief (y) and time (x)

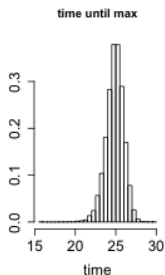
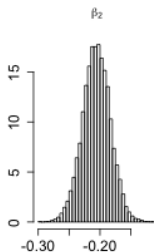
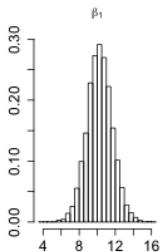
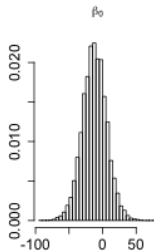
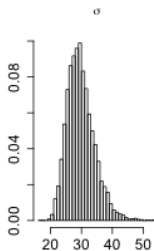
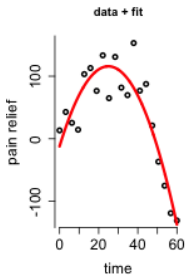
$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \varepsilon.$$

- At what time x_{max} is there **maximal pain relief**?

$$x_{max} = -\beta_1 / 2\beta_2$$

- Posterior distribution of x_{max} can be obtained by change of variable. Cauchy-like.
- Easy to obtain marginal posterior $p(x_{max} | \mathbf{y}, \mathbf{X})$ by **simulation**:
 - ▶ Simulate N coefficient vectors from the posterior $\beta, \sigma^2 | \mathbf{y}, \mathbf{X}$
 - ▶ For each simulated β , compute $x_{max} = -\beta_1 / 2\beta_2$.
 - ▶ Plot a histogram. Converges to $p(x_{max} | \mathbf{y}, \mathbf{X})$ as $N \rightarrow \infty$.

Finding the time for maximum



Bayes is easy to use

- Substantially more complex models can be analyzed by
 - ▶ **Markov Chain Monte Carlo** (MCMC) simulation
 - ▶ **Hamiltonian Monte Carlo** (HMC) simulation
 - ▶ **Variational inference** optimization
- Ongoing research on making Bayes more scalable to large data.
My own contributions: <https://mattiasvillani.com/research>
- Probabilistic programming languages (**Stan**) makes Bayes easy.
- Bayesian Learning course at SU:
<https://github.com/mattiasvillani/BayesLearnCourse>